## 6 Online Appendix (Not for Publication)

## 6.1 Existence of Optimal Deterministic Mechanism

In this appendix, we establish that an optimal deterministic mechanism exists.

**Proposition 3** There exists an optimal deterministic mechanism.

**Proof.** It suffices to show that for any  $\{r_1, ..., r_L\} \subset \{1, ..., T\}$  there exists an optimal mechanism in which the principal offers the stopping plan where the number of failures born by agents is described by the set  $\{r_1, ..., r_L\}$ . Notice that an incentive-compatible mechanism  $\mathcal{M}$  leads to L mappings  $\theta \to (a_i(\theta), b_i(\theta))$  for  $i \in \{1, ..., L\}$ , where  $a_l(\theta)$  stands for for the firm's lump-sum payment and  $b_l(\theta)$  for the success bonus of a type  $\tau = (\theta, c)$  who chooses a stopping plan with tolerance for  $r_l$  failures. Let  $\pi^*$  be the supremum of all payoffs obtained by mechanisms in which the principal offers the action plans  $\{r_1, ..., r_L\}$ . We will show that there exists a mechanism  $\mathcal{M}^*$  that yields the payoff  $\pi^*$  to the principal. For each  $n \in \mathbb{N}$ , take a sequence of mechanisms  $\mathcal{M}_n$  yielding a payoff to the principal at least as large as  $\pi^* - n^{-1}$ . The mechanism  $\mathcal{M}_n$  leads to the mappings  $(a_i^n(\theta), b_i^n(\theta))$  for  $i \in \{1, ..., L\}$ and to the threshold curves  $v_{r_l}^n(\theta)$ . Define  $\theta_{r_l}^n \equiv \inf \{\theta' : v_{r_l}^n(\theta') > 0\}$  for  $l \in \{1, ..., L\}$ . We will use the mappings  $(a_i^n(\theta), b_i^n(\theta))$  ( $i \in \{1, ..., L\}, n \in \mathbb{N}$ ) to construct our mechanism  $\mathcal{M}^*$ .

It is easy to show that we can restrict attention to mechanisms for which  $b_i^n(\theta) \ge 0$  for all  $\theta$ . Next, we claim that for each  $m \in \mathbb{N}$  and  $i \in \{1, ..., L\}$  we have

$$\sup_{n} \max_{\theta \in [0, 1-m^{-1}]} b_i^n(\theta) < \infty.$$
<sup>(29)</sup>

Assume towards a contradiction that there is  $m \in \mathbb{N}$  and a subsequence for which

$$\sup_{n} \max_{\theta \in [0, 1-m^{-1}]} b_i^n(\theta) = \infty.$$

Notice that the payoff of any type  $\theta' \in \left[1 - \frac{1}{2m}, 1\right]$  from choosing an allocation designed for a type  $\theta'' \in \left[0, 1 - \frac{1}{m}\right]$  is at least

$$a_i^n(\theta'') + b_i^n(\theta'') \cdot D_i(\theta') - \left(\sum_{t=1}^T \delta^{t-1}\right) \cdot \bar{c},$$

where  $D_i(\theta') \equiv \left[\sum_{t=1}^{r_i} \delta^{t-1} \lambda + (1 - (1 - \lambda)^{r_i}) \cdot \sum_{t=r_i+1}^T \delta^{t-1} \lambda\right] \cdot \theta'$ . Since  $a_i^n(\theta'') + b_i^n(\theta'') \cdot D_i(\theta'') \ge 0$  the expression above is at least as large as

$$b_i^n(\theta'') \cdot \left(D_i(\theta') - D_i(\theta'')\right) - \left(\sum_{t=1}^T \delta^{t-1}\lambda\right) \cdot \bar{c}$$
(30)

Hence, we conclude that (30) diverges to  $\infty$ . Consequently, so does the payoff of all types  $\theta' \in [1 - \frac{1}{2m}, 1]$ , which automatically implies that the principal obtains a negative payoff whenever n is large enough.

We will construct a contract  $(a_i(\theta), b_i(\theta)) = \{(a_i(\theta), b_i(\theta))\}_{i=1}^L$  from the sequence of contracts  $\{(a_i^n(\theta), b_i^n(\theta))\}_{i=1}^L$ . Notice that  $b_i^n : [\theta_{r_i}^n, 1) \to \mathbb{R}$  is increasing, while  $a_i^n : [\theta_{r_i}^n, 1) \to \mathbb{R}$  is decreasing. Notice that we may extend  $(a_i^n(\theta), b_i^n(\theta))$  to [0, 1) by letting  $(a_i^n(\theta), b_i^n(\theta)) = (a_i^n(\theta_{r_i}^n), b_i^n(\theta_{r_i}^n))$  for all  $\theta \leq \theta_{r_i}^n$  (notice that zero-measure sets have no impact on payoffs). Notice that (29) imply that  $b_i^n(\theta)$  is monotonic and uniformly bounded over the interval  $[0, 1 - m^{-1}]$  (for each m) and, thus, Helly's First Theorem (Theorem 6.1.18 in Kannan and Krueger 1996) asserts that there exists a subsequence  $b_i^{n_m}(\theta)$  which converges (a.e.) over  $[0, 1 - m^{-1}]$ . This property is also true for  $[0, 1 - (m + z)^{-1}]$  for all  $z \in \mathbb{N}$ , and hence we can find a subsequence of  $b_i^{n_m}(\theta)$ , call it  $b_i^{n_{m+1}}(\theta)$ , which converges over  $[0, 1 - (m + 1)^{-1}]$ . Proceeding inductively (by a diagonal argument) we obtain a subsequence of  $b_i^n(\theta)$ , call it  $b_i^n(\theta)$ , and an increasing function  $b_i(\theta)$  such that  $b_i^n(\theta) \to b_i(\theta)$  for almost all  $\theta \in [0, 1)$ . Since  $(a_i^n(\theta))$  is decreasing the same argument implies that we may take a subsequence  $(a_i^n(\theta))$  of  $(a_i^n(\theta))$  and a function  $a_i(\theta)$  such that  $a_i^n(\theta) \to a_i(\theta)$  for almost all  $\theta \in [0, 1)$ .

Proceeding analogously for all  $i \in \{1, ..., L\}$ , we obtain  $\{(a_1(\theta), b_1(\theta)), ..., (a_L(\theta), b_L(\theta))\}$ . We must show that

$$\pi^* = \int_0^1 \Pi(a(\theta), b(\theta)) f(\theta) d\theta.$$
(31)

Let  $S \equiv T\Delta$  and notice that  $\Pi(a(\theta), b(\theta)) - S \leq 0$  for all  $\theta$ . Let  $\mathbf{1}_{[\theta \leq 1-m^{-1}]}$  be the indicator function for  $\theta \leq 1 - m^{-1}$  and define  $g^m$  by

$$g^{m}(\theta) \equiv \left(\Pi(a(\theta), b(\theta)) - S\right) \cdot f(\theta) \cdot \mathbf{1}_{[\theta \le 1 - m^{-1}]}(\theta).$$

Notice that  $g_m$  is a decreasing sequence of nonpositive functions. Hence by the Lebesgue's monotone convergence theorem:

$$\int_0^1 \left( \Pi(a(\theta), b(\theta)) - S \right) f(\theta) d\theta = \lim_m \int_0^1 g_m(\theta) d\theta.$$
(32)

We claim that  $\int_0^1 (\Pi(a(\theta), b(\theta)) - S) f(\theta) d\theta > -\infty$ . Assume towards a contradiction that

$$\int_0^1 \left( \Pi(a(\theta), b(\theta)) - S \right) f(\theta) d\theta = -\infty.$$

In this case we can find  $\bar{m} \in \mathbb{N}$  such that  $\int_0^1 g_{\bar{m}}(\theta) d\theta < -4S$  and hence

$$\int_{0}^{1} \Pi(a(\theta), b(\theta)) \cdot f(\theta) \cdot \mathbf{1}_{[\theta \le 1 - \bar{m}^{-1}]}(\theta) \, d\theta < -3S$$

Thus we can find  $n^* \in \mathbb{N}$  such that  $n > n^*$  implies

$$\int_{0}^{1} \Pi(a^{n}(\theta), b^{n}(\theta)) \cdot f(\theta) \cdot \mathbf{1}_{[\theta \le 1 - \bar{m}^{-1}]}(\theta) \, d\theta < -2S$$
(33)

Since  $\int_0^1 \Pi(a^n(\theta), b^n(\theta)) \cdot f(\theta) \cdot \mathbf{1}_{[\theta > 1 - \bar{m}^{-1}]}(\theta) \, d\theta < S$ , (33) implies

$$\int_{0}^{1} \Pi(a^{n}(\theta), b^{n}(\theta)) \cdot f(\theta)(\theta) \, d\theta < -S, \tag{34}$$

which contradicts the assumption that  $\int \Pi(a^n(\theta), b^n(\theta)) \cdot f(\theta)(\theta) d\theta > \pi^* - \frac{1}{n} \geq -\frac{1}{n}$  whenever  $n > S^{-1}$ . Thus we have  $\int_0^1 \Pi(a(\theta), b(\theta)) f(\theta) d\theta > -\infty$ .

Take  $\varepsilon > 0$ . We must show that

$$\int_{0}^{1} \Pi(a(\theta), b(\theta)) f(\theta) d\theta \ge \pi^* - \varepsilon$$
(35)

to establish (31). Since  $\int_0^1 \Pi(a(\theta), b(\theta)) f(\theta) d\theta > -\infty$ , (32) implies that there is  $m_1 \in \mathbb{N}$  such that  $n > m_1$  implies

$$\int_{1-m_1^{-1}}^{1} \Pi(a(\theta), b(\theta)) f(\theta) d\theta > -\frac{\varepsilon}{4}.$$

Notice also that

$$\int_0^1 \Pi(a^n(\theta), b^n(\theta)) \cdot f(\theta) \cdot \mathbf{1}_{[\theta > 1 - m^{-1}]}(\theta) \, d\theta < S \cdot \left(1 - F(1 - m^{-1})\right)$$

Thus we take  $m_2 \ge m_1$  such that  $S \cdot (1 - F(1 - m_2^{-1})) < \frac{\varepsilon}{4}$  and  $n^* \in \mathbb{N}$  such that  $n^{*-1} < -\frac{\varepsilon}{4}$ . Take  $n^{**} > n^*$  such that

$$\left| \int_{0}^{1-m_{2}^{-1}} \Pi(a^{n^{**}}(\theta), b^{n^{**}}(\theta)) f(\theta) d\theta - \int_{0}^{1-m_{2}^{-1}} \Pi(a(\theta), b(\theta)) f(\theta) d\theta \right| < \frac{\varepsilon}{4}.$$

We have

$$\begin{split} & \int_0^1 \Pi(a(\theta), b(\theta)) f(\theta) d\theta \\ \geq & \int_0^{1-m_2^{-1}} \Pi(a(\theta), b(\theta)) f(\theta) d\theta - \frac{\varepsilon}{4} \\ \geq & \int_0^{1-m_2^{-1}} \Pi(a^{n^{**}}(\theta), b^{n^{**}}(\theta)) f(\theta) d\theta - \frac{\varepsilon}{2} \\ \geq & \int_0^1 \Pi(a^{n^{**}}(\theta), b^{n^{**}}(\theta)) f(\theta) d\theta - \frac{3\varepsilon}{4} \\ \geq & \pi^* - n^{**-1} - \frac{3\varepsilon}{4} \\ > & \pi^* - \varepsilon, \end{split}$$

which establishes (35) and completes the proof.

## 6.2 Correlation

In this appendix, we show that the result on the irrelevance of expertise regarding project quality (Proposition 1) does not rely on the independence between  $\theta$  and c. First, we construct and example

where expertise is irrelevant even though  $\theta$  and c are correlated. Second, we specialize the model to T = 1, and derive sufficient conditions for expertise to be irrelevant (therefore extending Proposition 1 to environments with correlation). Third, we show that Proposition 1 is robust to small perturbations away from independence (in the form of mixture distributions).

Consider the following example.

**Example A 1** [Conditional Distribution is Mirrored Generalized Pareto] Let the distribution of the safe project payoff conditional on the quality of the risky project be

$$H(c|\theta) = \left(\frac{c-\underline{c}}{\overline{c}-\underline{c}}\right)^{\eta\theta}, \quad where \quad \eta > 0.$$

Notice that the conditional distribution  $H(\cdot|\theta)$  increases in the sense of first-order stochastic dominance as  $\theta$  increases, in which case  $\theta$  and c are positively correlated. In this case, the conditional reverse hazard rate is  $\gamma(c|\theta) \equiv \frac{H(c|\theta)}{h(c|\theta)} = \frac{c-c}{\eta\theta}$ . Consider the action plan  $\phi^e$  as described Lemma 3, after replacing the unconditional reverse hazard rate  $\gamma(c)$  by its conditional counterpart  $\gamma(c|\theta)$ , and let  $P^e(\tau)$  be the expected payments induced by the expert-investor optimal mechanism (formula 9).

Following the same reasoning as in the proof of Proposition 1, it follows that the action plan  $\phi^e$  is implementable by a menu of linear contracts with lump-sum payments and success bonuses:

$$a^*(\tau) = \frac{\Phi_{k^*(\tau)}(\theta)}{\eta + 1} \cdot (\underline{c} - \eta \theta K) \quad and \quad b^*(\tau) = \triangle \cdot \left(1 + \frac{1}{\eta \theta}\right)^{-1}.$$

To understand the implementability claim, note that ICS is satisfied by construction, ICR<sub>1</sub> holds as  $b^*(\tau)$  is increasing in  $\theta$ , and ICR<sub>2</sub> holds by the same argument as in the proof of Proposition 1. Because, by construction, expected payments under two-dimensional asymmetric information equal  $P^e(\tau)$ , it follows that expertise (about  $\theta$ ) is irrelevant for payoffs.

The example above presents a parametric case exhibiting correlation where the result and proof technique of Proposition 1 readily apply. In what follows, we will derive sufficient conditions for expertise to be irrelevant. For tractability, we will assume that T = 1. As in the example above, let  $H(c|\theta)$  denote the conditional cumulative distribution of c given  $\theta$ , let  $h(c|\theta) = \frac{\partial H}{\partial c}(c|\theta)$  denote its density, and let  $\gamma(c|\theta) \equiv \frac{H(c|\theta)}{h(c|\theta)}$  denote its associated reverse hazard rate. As in the case of independence, we assume that  $H(c|\theta)$  is log-concave in c for each  $\theta$ , so that  $\frac{\partial \gamma}{\partial c}(c|\theta) \ge 0$ .

From the same argument as in the text, when there is symmetric information about  $\theta$ , the agent experiments if  $c \leq v(\theta)$ , where  $v(\theta)$  is the implicit solution of:

$$\theta \cdot \lambda \cdot \triangle - K = v(\theta) + \gamma \left( v(\theta) | \theta \right). \tag{36}$$

Differentiating (36), gives

$$v'(\theta) = \frac{\lambda \cdot \triangle - \frac{\partial \gamma}{\partial \theta} \left( v(\theta) | \theta \right)}{1 + \frac{\partial \gamma}{\partial c} \left( v(\theta) | \theta \right)}.$$
(37)

Notice that the log-concavity of  $H(c|\theta)$  guarantees that the denominator above is positive. When there is correlation, in order to guarantee that  $v'(\theta) \ge 0$ , we also need that  $\frac{\partial \gamma}{\partial \theta} \le \lambda \cdot \Delta$ , which we assume from now on. This condition limits by how much an increase in  $\theta$  shifts the conditional reverse hazard rate of c. In intuitive terms, it requires that the correlation between  $\theta$  and c is not too negative.

As in the case of independence, the expert-investor optimal mechanism is incentive compatible when  $\theta$  is private information if and only if  $v(\theta)$  is convex. When c and  $\theta$  are independent, we have  $\frac{\partial \gamma}{\partial \theta} = 0$  and, therefore, v is convex if and only if  $\frac{\partial \gamma}{\partial c}$  is decreasing, i.e., the reverse hazard rate is weakly concave (Condition C).

For the general case, differentiate (37) again to obtain:

$$v''(\theta) = -\frac{\frac{\partial^2 \gamma}{\partial \theta^2} \left( v(\theta) | \theta \right) + v'(\theta) \left[ \frac{\partial^2 \gamma}{\partial c^2} \left( v(\theta) | \theta \right) v'(\theta) + 2 \frac{\partial^2 \gamma}{\partial c \partial \theta} \left( v(\theta) | \theta \right) \right]}{1 + \frac{\partial \gamma}{\partial c} \left( v(\theta) | \theta \right)}$$

Therefore, the following are sufficient conditions for expertise to be irrelevant:  $\frac{\partial^2 \gamma}{\partial c^2}, \frac{\partial^2 \gamma}{\partial \theta^2}, \frac{\partial^2 \gamma}{\partial \theta \partial c} \leq 0$ . As such, in the case of correlation, expertise is irrelevant whenever  $\gamma(c|\theta)$  is concave in each of its arguments and submodular.

Next, we show that our main result is robust to small perturbations away from independence.

Let H(c) be a log-concave distribution with a strictly concave reverse hazard rate. Let  $\Upsilon(c, \theta)$  be a joint distribution with a smooth density that is bounded away from zero in its support  $[\underline{c}, \overline{c}] \times [0, 1]$ .

Consider the mixture distribution

$$Q^{\alpha}(c,\theta) := \alpha H(c) + (1-\alpha)\Upsilon(c,\theta),$$

where  $0 \leq \alpha \leq 1$ . As  $\alpha$  approaches 1, this distribution converges to H. For each  $\theta$ , let  $c \to Q^{\alpha}(c \mid \theta)$  represent the marginal distribution associated with  $\Upsilon$ . Let  $q(c \mid \theta)$  denote its density and  $\gamma^{\alpha}(c, \theta) := \frac{Q^{\alpha}(c \mid \theta)}{q^{\alpha}(c \mid \theta)}$  denote its reverse hazard rate.

For each  $\alpha$ , let  $v(\theta, \alpha)$  denote the implicit solution of (36), and let

$$\vartheta\left(\theta,\alpha\right) \equiv -\frac{\frac{\partial^{2}\gamma^{\alpha}}{\partial\theta^{2}} + \frac{\lambda\Delta - \frac{\partial\gamma^{\alpha}}{\partial\theta}}{1 + \frac{\partial\gamma^{\alpha}}{\partial c}} \left[\frac{\partial^{2}\gamma^{\alpha}}{\partial c^{2}} \cdot \frac{\lambda\Delta - \frac{\partial\gamma^{\alpha}}{\partial\theta}}{1 + \frac{\partial\gamma^{\alpha}}{\partial c}} + 2\frac{\partial^{2}\gamma^{\alpha}}{\partial\theta\partial c}\right]}{1 + \frac{\partial\gamma^{\alpha}}{\partial c}},$$

where we omit the  $(v(\theta, \alpha), \theta)$  from all functions on the right-hand side for notational simplicity. As argued previously, expertise is irrelevant when  $\vartheta(\theta, \alpha) \ge 0$  for all  $\theta$  (for a fixed  $\alpha$ ). We claim that there exists  $\alpha^* \in (0, 1)$  such that  $\vartheta(\theta, \alpha) > 0$  for all  $\alpha > \alpha^*$ . Since  $(\theta, c) \to v(\theta, \alpha)$  is smooth, so is  $(\theta, \alpha) \to \vartheta(\theta, \alpha)$ . Moreover, by our assumption on H,  $\vartheta(\theta, 1) > 0$  for all  $\theta$ . The result then follows by uniform continuity.

## References

 Kannan, R. and Krueger, C. K. (1996): Advanced Analysis on the Real Line, Springer, New York.