

Complementary Proofs to “Careerist Judges”

Proof of Lemma 6: (i) $q_n(n, t^*)$ increases with t^* :

$$\begin{aligned}
 & \text{sign} \frac{\partial q_n(n, t^*)}{\partial t^*} \\
 &= \text{sign}(1 - q)qf(t^*)[-t^* \int_{t^*}^1 (1 - v)f(v)dv + (1 - t^*) \int_{t^*}^1 vf(v)dv] \\
 &= \text{sign} - t^* \int_{t^*}^1 f(v)dv + t^* \int_{t^*}^1 vf(v)dv + \int_{t^*}^1 vf(v)dv - t^* \int_{t^*}^1 vf(v)dv \\
 &= \text{sign}(1 - q)qf(t^*)[-t^* \int_{t^*}^1 f(v)dv + \int_{t^*}^1 vf(v)dv] > 0
 \end{aligned}$$

Similarly, I will show that $q_n(y, t^*)$ decreases with t^* :

$$\begin{aligned}
 & \text{sign} \frac{\partial q_n(s^* = y, t^*)}{\partial t^*} \\
 &= \text{sign}(1 - t^*)(\int_{.5}^1 (1 - v)f(v)dv + \int_{.5}^{t^*} vf(v)dv) - t^*(\int_{.5}^1 vf(v)dv + \int_{.5}^{t^*} (1 - v)f(v)dv) \\
 &= \text{sign}(1 - q)qf(t^*)(\int_{.5}^{t^*} (1 - 2t^*)f(v)dv + \int_{t^*}^1 (1 - t^* - v)f(v)dv) < 0. \square
 \end{aligned}$$

(ii) The proof of this is analogous to part (i). \square

(iii) Follows from (i) and (ii). \square

(iv) By the above claims, it is enough to show that $q_y(n, t^*) > q$, i.e., that:

$$\frac{q(\int_{.5}^1 vf(v)dv + \int_{.5}^{t^*} (1 - v)f(v)dv)}{q(\int_{.5}^1 vf(v)dv + \int_{.5}^{t^*} (1 - v)f(v)dv) + (1 - q)(\int_{.5}^1 (1 - v)f(v)dv + \int_{.5}^{t^*} vf(v)dv)} > q$$

which holds if

$$\begin{aligned}
 \int_{.5}^1 vf(v)dv + \int_{.5}^{t^*} (1 - v)f(v)dv &> \int_{.5}^1 (1 - v)f(v)dv + \int_{.5}^{t^*} vf(v)dv \iff \\
 \int_{t^*}^1 vf(v)dv &> \int_{t^*}^1 (1 - v)f(v)dv \iff \int_{t^*}^1 (2v - 1)f(v)dv > 0.
 \end{aligned}$$

The second part, i.e., that $q_n(s^*, t^*) > 1 - q$, has an analogous proof. \square

(v) To show that $q_y(n, t^*) > \Pr(w = y|q, n, t^*)$, we need to show that:

$$q_y(n, t^*) > \frac{q(1 - t^*)}{q(1 - t^*) + t^*(1 - q)}$$

which, after re-arranging, is analogous to:

$$t^*(\int_{.5}^1 vf(v)dv + \int_{.5}^{t^*} (1 - v)f(v)dv) > (1 - t^*)(\int_{.5}^1 (1 - v)f(v)dv + \int_{.5}^{t^*} vf(v)dv)$$

where the last inequality is satisfied using the proof in part (iv). To show that $q_y(y, t^*) > \Pr(w = y|q, y, t^*)$, we need to show that:

$$q_y(y, t^*) > \frac{qt^*}{qt^* + (1-t^*)(1-q)} \iff \frac{\int_{t^*}^1 vf(v)dv}{\int_{t^*}^1 (1-v)f(v)dv} > \frac{t^*}{1-t^*},$$

but

$$\int_{t^*}^1 vf(v)dv > t^* \text{ and } \int_{t^*}^1 (1-v)f(v)dv < 1-t^*,$$

hence the above is satisfied. We can use the analogous proof to show that $q_n(s^*, t^*) > \Pr(w = n|q, s^*, t^*)$. \square

(vi) Note that $q_y(y, \frac{1}{2}) > q_n(y, \frac{1}{2})$, and by the above claims this holds for all $t^* > \frac{1}{2}$ and $s^* = y$. On the other hand, $q_y(n, t^* \rightarrow 1) \rightarrow q$ and $q_n(n, t^* \rightarrow 1) \rightarrow 1$. Since $q_n(n, t^*)$ increases with t^* and $q_y(n, t^*)$ decreases with t^* , there exists a unique $\tilde{t}(q)$ such that $q_y(n, \tilde{t}(q)) = q_n(n, \tilde{t}(q))$ and for all $t^* < (>) \tilde{t}(q)$, $q_y(n, t^*) > (<) q_n(n, t^*)$. With the uniform distribution, i.e., $f(v) = 2$, then $q_y(n, q) = \frac{2-q}{3-2q} \geq \frac{1+q}{1+2q} = q_n(n, q)$ for all $q \geq .5$ which implies that $\tilde{t}(q) > q$. \square

(vii) To see this, recall that:

$$\tilde{p}(y) = q_y^2 + (1 - q_y) \Pr(w = y|q, s, t),$$

hence:

$$\frac{\partial \tilde{p}(y)}{\partial t} \Big|_{s^c=n} = (2q_y - \Pr(w = y|q, n, t)) \frac{\partial q_y}{\partial t} + (1 - q_y) \frac{\partial \Pr(w = y|q, n, t)}{\partial t}$$

but when $s^c = n$, $\frac{\partial q_y}{\partial t} < 0$. Also, $\frac{\partial \Pr(w=y|q, n, t)}{\partial t}$ and $2q_y - \Pr(w = y|q, s, t) > 0$ by the above. Similarly,

$$\frac{\partial \tilde{p}(n)}{\partial t} = (2q_n - \Pr(w = n|q, n, t)) \frac{\partial q_n}{\partial t} + (1 - q_n) \frac{\partial \Pr(w = n|q, n, t)}{\partial t} > 0.$$

An analogous analysis holds for the derivatives w.r.t. q .

This completes the proof of Lemma 6. \blacksquare

Proof of Proposition 1, Step 2 (Uniqueness). I show a sufficient condition for uniqueness, i.e., that at the equilibrium value of t^e , whenever $\frac{\partial q_y(n, t) - \theta\beta(n, t)}{\partial t q_n(n, t) + \theta\beta(n, t)} < 0$, then:

$$\left| \frac{\partial \Pr(w = y|q, n, t)}{\partial t \Pr(w = n|q, n, t)} \right| > \left| \frac{\partial q_y(n, t) - \theta\beta(n, t)}{\partial t q_n(n, t) + \theta\beta(n, t)} \right|^{.1}$$

Consider first $\frac{\Pr(w=y|q, n, t)}{\Pr(w=n|q, n, t)} = \frac{q(1-t)}{t(1-q)}$. Then:

$$\left| \frac{\partial \Pr(w = y|q, n, t)}{\partial t \Pr(w = n|q, n, t)} \right| = \frac{q}{(1-q)t^2}$$

Now consider

$$\begin{aligned} & \left| \frac{\partial q_y(n, t) - \theta(\beta(n, t))}{\partial t q_n(n, t) + \theta\beta(n, t)} \right| \\ &= \frac{1}{q_n + \theta\beta} \left[\frac{\partial q_y}{\partial t} (-1 + \theta \frac{\partial \beta}{\partial q_y} + \theta \frac{\partial \beta}{\partial q_y} \frac{q_y - \theta\beta}{q_n + \theta\beta}) + \frac{\partial q_n}{\partial t} \left((1 + \theta \frac{\partial \beta}{\partial q_n}) \frac{q_y - \theta\beta}{q_n + \theta\beta} + \theta \frac{\partial \beta}{\partial q_n} \right) \right] \end{aligned}$$

but since

$$\frac{\partial \beta}{\partial q_y} > 0, \frac{\partial \beta}{\partial q_n} < 0, \frac{\partial q_n}{\partial t} > 0 \text{ and } \frac{\partial q_y}{\partial t} < 0,$$

it is enough to show that

$$\frac{q}{(1-q)t^2} > \frac{1}{q_n} \left[-\frac{\partial q_y}{\partial t} + \frac{\partial q_n}{\partial t} \frac{q_y - \theta\beta}{q_n + \theta\beta} \right]$$

Plugging in the equilibrium condition and the expressions for the derivatives, the right-hand-side becomes:

$$\frac{1}{q_n} \left[\frac{2q_y(1-q_y)}{(2-t)t} + \frac{q}{(1+t)} \frac{2q_n(1-q_n)}{t(1-q)} \right]$$

Let $q_x \in \{q_y, q_n\}$ such that $q_x(1-q_x) = \max\{q_y(1-q_y), q_n(1-q_n)\}$. It is therefore sufficient to prove that:

$$\frac{q}{t} > \frac{2q_x(1-q_x)}{q_n} \left[\frac{(1+t)(1-q) + q(2-t)}{(2-t)(1+t)} \right]$$

But the above equation holds both when $q_x = q_n$ and when $q_x = q_y$. ■

Proof of Proposition 1, Step 4 ($\frac{\partial t^e(q, \theta)}{\partial q} > 0$).

By total differentiation of the equilibrium condition:

$$\frac{\partial t}{\partial q} \Big|_{t=t^e} = \frac{\frac{\partial}{\partial q} \frac{q_y(n, t) - \theta\beta(n, t)}{q_n(n, t) + \theta\beta(n, t)} - \frac{\partial}{\partial q} \frac{\Pr(w=y|q, n, t)}{\Pr(w=n|q, n, t)}}{\frac{\partial}{\partial t} \frac{\Pr(w=y|q, n, t)}{\Pr(w=n|q, n, t)} - \frac{\partial}{\partial t} \frac{q_y(n, t) - \theta\beta(n, t)}{q_n(n, t) + \theta\beta(n, t)}} \Big|_{t=t^e}$$

I show that when $\frac{\Pr(w=y|q, n, t)}{\Pr(w=n|q, n, t)} = \frac{q_y(n, t) - \theta\beta(n, t)}{q_n(n, t) + \theta\beta(n, t)}$,

$$\frac{\partial}{\partial q} \frac{\Pr(w=y|q, n, t)}{\Pr(w=n|q, n, t)} > \frac{\partial}{\partial q} \frac{q_y(n, t) - \theta\beta(n, t)}{q_n(n, t) + \theta\beta(n, t)}$$

which, along with step 2 of the Proposition, implies that $\frac{\partial t}{\partial q} \Big|_{t=t^e} > 0$. As in step 2, it is enough to

show the inequality for $\theta = 0$, i.e., we have to show that:

$$\begin{aligned} \frac{(1-t)}{t(1-q)^2} &> \frac{1}{q_n} \left[\frac{\partial q_y}{\partial q} - \frac{\partial q_n}{\partial q} \frac{q(1-t)}{t(1-q)} \right] \iff \\ \frac{(1-t)}{t(1-q)^2} &> \frac{1}{q_n} \left[\frac{q_y(1-q_y)}{q(1-q)} + \frac{q_n(1-q_n)(1-t)}{t(1-q)^2} \right] \end{aligned}$$

but since $q_y > q_n$ in equilibrium, and for all q and t , $q_y > 1 - q_n$, it is sufficient to show that:

$$\frac{(1-t)}{t(1-q)^2} > \frac{q_n(1-q_n)}{q_n} \left[\frac{1}{q(1-q)} + \frac{(1-t)}{t(1-q)^2} \right] \iff$$

$$\frac{q(1-t)}{t(1-q) + q(1-t)} > \frac{q(1-t)}{(1-q)(1+t) + q(1-t)}$$

which holds for all $t, q \in [.5, 1]$. This implies that $\frac{\partial t}{\partial q}|_{t=t^e} > 0$. ■

Proof of Proposition 2, Step 2 ($t^c(q) < \hat{t}$ for all q) :

Let $s^c = n$ and $t = t^c$. I will show that there is a unique $\hat{t} < 1$ satisfying $\tau_{\hat{t}}(y, y, \sigma^c) = \tau_{\hat{t}}(n, y, \sigma^c)$, and that for all $t > \hat{t}$, $\tau(y, y, \sigma^c) < \tau(n, y, \sigma^c)$. This implies that an equilibrium with $t > \hat{t}$ cannot exist, since then the expected utility from ruling n , an average over $\tau(n, y, \sigma^c)$ and $\tau(n, n, \sigma^c)$ where $\tau(n, n, \sigma^c) > \tau(n, y, \sigma^c)$, is higher than the expected utility from ruling y , an average over $\tau(y, n, \sigma^c)$ and $\tau(y, y, \sigma^c)$ where $\tau(y, y, \sigma^c) > \tau(y, n, \sigma^c)$.

The expression for $\tau(y, y, \sigma^c)|_{s^c=n}$ is

$$\tau(y, y, \sigma^c)|_{s^c=n} = \frac{\int_{.5}^1 t^2 f(t) dt + \int_{.5}^{t^c} t(1-t) f(t) dt}{\int_{.5}^1 t f(t) dt + \int_{.5}^{t^c} (1-t) f(t) dt}$$

Taking the derivative of $\tau(y, y, \sigma^c)|_{s^c=n}$ w.r.t t^c , it is

$$\frac{d\tau(y, y, \sigma^c)}{dt^c}|_{s^c=n} = \frac{(1-t^c)f(t^c)(t^c - \tau(y, y, \sigma^c))}{(\int_{.5}^1 t f(t) dt + \int_{.5}^{t^c} (1-t) f(t) dt)^2}$$

Therefore, $\tau(y, y, \sigma^c)$ is a monotonously decreasing function as long as $t^c < \tau(y, y, \sigma^c)$ and a monotonously increasing function when $t^c > \tau(y, y, \sigma^c)$. When $t^c \rightarrow .5$, $t^c < \tau(y, y, \sigma^c)$ and when $t^c \rightarrow 1$, $\tau(y, y, \sigma^c) < t^c$. Therefore, there exists t' such that $t' = \tau_{t'}(y, y, \sigma^c)$. Moreover, t' is unique, since when $t^c > \tau(y, y, \sigma^c)$, $\frac{d\tau(y, y, \sigma^c)}{dt^c} < 1$. Thus, for all $t^c < (>)t'$, $t^c < (>)\tau(y, y, \sigma^c)$ and $\frac{d\tau(y, y, \sigma^c)}{dt^c} < (>)0$.

On the other hand, $\tau(n, y, \sigma^c)$ is an average over t for $t > t^c$ and thus increases with t^c for all $t^c > .5$. Also, since only values of $t > t^c$ are included in the computation of $\tau(n, y, \sigma^c)$, then $\tau(n, y, \sigma^c) > t^c$ for all t^c . By the above, when $t^c \rightarrow 1$, $\tau(n, y, \sigma^c) > t^c > \tau(y, y, \sigma^c)$. When $t^c = .5$, by Lemma 3, $\tau(y, y, \sigma^c) = \tau(n, n, \sigma^c) > \tau(n, y, \sigma^c)$. Then, there must exist some $\hat{t} \in (.5, 1)$ satisfying $\tau_{\hat{t}}(y, y, \sigma^c) = \tau_{\hat{t}}(n, y, \sigma^c)$. Moreover, it must be that $\hat{t} < t'$ because for all $t^c > t'$, $\tau(n, y, \sigma^c) > t^c > \tau(y, y, \sigma^c)$. Because $\tau(y, y, \sigma^c)$ decreases monotonously for $t^c < \hat{t}$ and $\tau(n, y, \sigma^c)$ increases monotonously in t^c , \hat{t} is unique. ■

Proof of Proposition 4, Part (ii). We have to show that $t^f(q) < t^c(q)$. $t^f(q)$ solves:

$$\Pr(w = y|q, n, t^f(q))\Gamma_y = \Pr(w = n|q, n, t^f(q))\Gamma_n + \Gamma \quad (1)$$

where $\Gamma_y = \tau(y, y, \sigma^f) - \tau(y, n, \sigma^f)$, $\Gamma_n = \tau(n, n, \sigma^f) - \tau(n, y, \sigma^f)$ and $\Gamma = \tau(n, y, \sigma^f) - \tau(y, n, \sigma^f)$. I will show that at $t^f(q)$,

$$\tilde{p}(y)\Gamma_y > \tilde{p}(n)\Gamma_n + \Gamma \quad (2)$$

for $\tilde{p}(d) = (1 - q_d) \Pr(w = d|q, n, t^f(q)) + q_d^2$, which implies that at $t^f(q)$, the utility from y is higher than the utility from n if appeals are allowed, meaning that the equilibrium solution $t^c(q)$ must admit $t^c(q) > t^f(q)$.

Plugging (1) into (2), we have to show that:

$$\begin{aligned} q_y(q_y - \Pr(w = y|q, n, t^f(q)))\Gamma_y > q_n(q_n - \Pr(w = n|q, n, t^f(q)))\Gamma_n &\Leftrightarrow \\ \frac{\Gamma_y}{\Gamma_n} > \frac{q_n(q_n - \Pr(w = n|q, n, t^f(q)))}{q_y(q_y - \Pr(w = y|q, n, t^f(q)))} \end{aligned}$$

However, I now show that for all values of t , $\frac{\Gamma_y}{\Gamma_n} > 1$, whereas for all values of t , $\frac{q_n(q_n - \Pr(w = n|q, n, t^f(q)))}{q_y(q_y - \Pr(w = y|q, n, t^f(q)))} < 1$, which implies the desired result.

To see that $\frac{\Gamma_y}{\Gamma_n} > 1$, I calculate the reputations for a cutoff point t from each action and state of the world:

$$\begin{aligned} \Gamma_y &= \tau(y, y, \sigma) - \tau(y, n, \sigma) \\ &= \frac{\int_{.5}^1 2v^2 dv + \int_{.5}^t 2v(1-v)dv}{\int_{.5}^1 2v dv + \int_{.5}^t 2(1-v)dv} - \frac{\int_{.5}^1 2v(1-v)dv + \int_{.5}^t 2v^2 dv}{\int_{.5}^1 2(1-v)dv + \int_{.5}^t 2v dv} \\ &= \frac{\frac{1}{2}t - \frac{1}{3}t^3 - \frac{1}{6}}{t^2(2-t)} \end{aligned}$$

and similarly

$$\Gamma_n = \tau(n, n, \sigma) - \tau(n, y, \sigma) = \frac{-t + t^2 + \frac{1}{3} - \frac{1}{3}t^3}{(1+t)(1-t)^2}$$

and therefore

$$\Gamma_y > \Gamma_n \iff \frac{1 - 4t^4 + 4t^3 - t^2 + 2t - 1}{6(t+1)t^2(2-t)} > 0$$

which holds for all $t > \frac{1}{2}$.

To see that $\frac{q_n(q_n - \Pr(w = n|q, n, t))}{q_y(q_y - \Pr(w = y|q, n, t))} < 1$ for all t , I simplify the expression and find that this holds iff:

$$\frac{(1-q)(1+t)(1-t)}{t^2(1-2q)^2 + 1 + 2t(1-2q)} < \frac{q(2-t)t}{t^2(1-2q)^2 + 4q^2 + 4qt(1-2q)}$$

which holds for all $t \geq \frac{1}{2}$ and $q \geq \frac{1}{2}$. This completes the proof. ■