MEI MT Problem Set 1

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Introduction

Your class teacher:

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- http://personal.lse.ac.uk/carayolt/EC402.htm
- Office hours: Tuesdays 11.30 to 12.30 in S684.
- ➤ 3 class groups on Fridays: 9 to 10 in K105, 3 to 4 in K105 and 5 to 6 in D211.

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My slides will be available on my website after classes.

What is a random variable?



► (a) and (b) basically same question, but (a) discrete and (b) continuous: given the distribution of r.v. X, we want the distribution of Y = f(X) = X² + 2.

The way to answer is very different.

Question 2 (a)

$$X = \begin{cases} -1 \text{ with probability } \frac{1}{3} \\ 0 \text{ with probability } \frac{1}{3} \\ 1 \text{ with probability } \frac{1}{3} \end{cases}$$

We can conclude directly, as this maps, through function f, to:

$$Y = f(X) = \begin{cases} 3 \text{ with probability } \frac{2}{3} \\ 2 \text{ with probability } \frac{1}{3} \end{cases}$$

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Question 2 (b)

- ➤ X uniformly distributed on [-1, +1]; i.e. its pdf is .5 on that interval, 0 elsewhere.
- ▶ Remember September course: we had a formula to derive the pdf of a *monotonic* function of a random variable. Here, $f(X) = X^2 + 2$ is *NOT* monotonic on [-1, +1] so we can't apply this directly.
- Let us think in terms of cdf. The cdf of X is:

$$F_X(x) = \begin{cases} 0 \text{ if } x \le -1 \\ \frac{x+1}{2} \text{ if } -1 \le x \le 1 \\ 1 \text{ if } 1 \le x \end{cases}$$

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Question 2 (b)

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$$F_{Y}(y) = P(Y \le y) = P(X^{2} + 2 < y) = P(X^{2} < y - 2)$$

= $P(-\sqrt{y - 2} < X < \sqrt{y - 2})$
= $F_{X}(\sqrt{y - 2}) - F_{X}(-\sqrt{y - 2})$
= $2F_{X}(\sqrt{y - 2}) - 1$
= $\begin{cases} 0 \text{ if } y \le 2 \\ \sqrt{y - 2} \text{ if } 2 \le y \le 3 \\ 1 \text{ if } 3 \le y \end{cases}$ (1)

(From there, can differentiate to get pdf)

Linearity of expectations.

$$E(aX + bY) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} (ax + by) f_{X,Y}(x, Y) dx dy$$

= $a \int_{-\infty}^{+\infty} x \left(\int_{-\infty}^{+\infty} f_{X,Y}(x, y) dy \right) dx$
+ $b \int_{-\infty}^{+\infty} y \left(\int_{-\infty}^{+\infty} f_{X,Y}(x, y) dx \right) dy$
= $a \int_{-\infty}^{+\infty} x f_X(x) dx + b \int_{-\infty}^{+\infty} y f_Y(y) dy$

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- Important question about the relationship between independence, mean-independence and uncorrelatedness. See also p.99 of your handout.
- (a) Independence stronger than mean-independence.
- ► (b) Mean-independence stronger than uncorrelatedness.
- Also noteworthy:
 - Mean-independence of X on Y is not the same thing as mean-independence of Y on X.
 - X mean-independent on Y and Y mean-independent on X does NOT imply X and Y independent (can find counterexample)
 - A subtle point. The statement on top of p.100 is only correct if X and Y are *jointly* normal, which is stronger than saying that they are both normal. More on this next week.

Question 4 (a)

$$E(X|Y) = \int_{-\infty}^{+\infty} x f_{X|Y}(x|Y) dx = \int_{-\infty}^{+\infty} x f_X(x) dx \text{ by independence}$$
$$= E(X)$$

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Question 4 (b)

$$cov(X, Y) = E(XY) - E(X)E(Y)$$

= $E(E(XY|Y)) - E(X)E(Y)$
= $E(YE(X|Y)) - E(X)E(Y)$
= $E(YE(X)) - E(X)E(Y)$
= $E(X)E(Y) - E(X)E(Y) = 0$

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Definition

 $X_1, ..., X_n$ are jointly normal (i.e., the vector $(X_1, X_2, ..., X_n)$ is gaussian) if either of these (equivalent) propositions is satisfied:

Any linear combination of the X_i's is normal (in particular, each X_i is normal-but that is NOT sufficient)

► The joint pdf of
$$X_1, ..., X_n$$
 is
 $f(x_1, ..., x_n) = \frac{1}{(2\pi)^{\frac{n}{2}}\sqrt{det(\Sigma)}} e^{-\frac{1}{2}((x-m)'\Sigma^{-1}(x-m))}$, where *m* is
the expected value of the gaussian vector and Σ is its
variance-covariance matrix.

- So here, by definition we know that Z = 5 + 2Y − X is normally distributed, as it is a linear combination of the elements of a gaussian vector. So all we have to do is find its mean and variance.
- $E(Z) = 5 + 2E(Y) E(X) = 5 + 2\mu_Y \mu_X$
- ► $Var(Z) = Var(2Y X) = Var(2Y) + Var(X) 2cov(2Y, X) = 4Var(Y) + Var(X) 4cov(X, Y) = 4\sigma_Y^2 + \sigma_X^2 4\sigma_{XY}$
- ► Hence we conclude $Z \sim \mathcal{N}(5 + 2\mu_Y \mu_X, 4\sigma_Y^2 + \sigma_X^2 4\sigma_{XY})$.

Question 6 and 7

- ▶ Question 6: sample
- Question 7: tests
- Let's talk about it quickly in class.

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- X_i : uniform distribution over $[\theta, 100]$.
- ▶ Method of moments: $\hat{\theta}_{MM}$ such that the sample counterpart of the first moment (e.g., sample mean) would coincide with the population mean for $\theta = \hat{\theta}_{MM}$.
- ► Here: sample mean is \bar{X} ; population mean is $\frac{100+\theta}{2}$. So $100 + \hat{\theta}_{MM} = 2\bar{X}$; hence $\hat{\theta}_{MM} = 2\bar{X} 100 = 2 \times 50 100 = 0$

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