

## Problem Set 2 – Solutions

### 1. Hypothesis test

1. Consider the  $MA(1)$  model

$$y_t = \varepsilon_t + \theta\varepsilon_{t-1}, \quad \varepsilon_t \sim iidN(0, \sigma^2), \quad t = 1, \dots, T, \quad \varepsilon_0 = 0$$

1. Derive the  $LM$  test of the null  $\theta = 0$
2. What are the advantages of the  $LM$  test over the  $LR$  test in this case?

**Answer.** *ML estimation of  $\theta$  is equivalent to non-linear least squares*

$$\min_{\theta} \sum_t \varepsilon_t(\theta)^2 \text{ where } \varepsilon_t(\theta) = y_t - \theta\varepsilon_{t-1}(\theta)$$

$$z_t = -\frac{\partial \varepsilon_t}{\partial \theta} = \varepsilon_{t-1} + \theta \frac{\partial \varepsilon_{t-1}}{\partial \theta}$$

$$FOC : \quad \sum_t z_t \varepsilon_t = 0.$$

*Quite complicated to compute.*

LM test. *Under the null,  $\theta = 0$  and  $y_t = \varepsilon_t$ . So  $z_t = \varepsilon_{t-1}$  and LM test is carried out by regressing  $\varepsilon_t$  on  $z_t$ , i.e. in this case by regressing  $y_t$  on  $y_{t-1}$  and*

---

\*© 2007 by Christian Julliard. This document may be reproduced for educational and research purposes, so long as the copies contain this notice and are retained for personal use or distributed free.

computing  $TR^2$ . Since this is a one variable linear regression,  $R^2$  is the square of the correlation coefficient, so

$$TR^2 = T \frac{(\sum y_t y_{t-1})^2}{(\sum y_t^2)(\sum y_{t-1}^2)} \sim \chi^2(1).$$

Therefore, the LM test is straightforward to compute in this case.

2. Consider the model

$$y_t = \beta_1 x_{1t} + \beta_2 \frac{(x_{2t} - \gamma)^{-2}}{2} + \varepsilon_t$$

where  $\varepsilon_t \sim iidN(0, \sigma^2)$  and the regressors  $x_1$  and  $x_2$  are process independent of the errors.

1. Obtain the log likelihood of the model and outline how you would obtain the NLE of the parameters  $\beta_1, \beta_2, \gamma, \sigma^2$ .
2. Construct a LM test of the null  $H_0 : \gamma = 0$ .
3. Why is the LM test easier to carry out than the Likelihood ratio test?

**Answer.**

$$\begin{aligned} \varepsilon_t &= y_t - \beta_1 x_{1t} - \beta_2 \frac{(x_{2t} - \gamma)^{-2}}{2} \\ \log L &= -\frac{T}{2} \log 2\pi - \frac{T}{2} \log \sigma^2 - \frac{1}{2\sigma^2} \sum \varepsilon_t^2. \end{aligned}$$

As this is nonlinear least squares maximize  $\log L$  wrt  $(\beta_1, \beta_2, \gamma)$  using derivations:

$$\begin{bmatrix} \frac{\partial \log L}{\partial \beta_1} \\ \frac{\partial \log L}{\partial \beta_2} \\ \frac{\partial \log L}{\partial \gamma} \end{bmatrix} = \frac{1}{\sigma^2} \sum z_t \varepsilon_t \text{ where } z_t = - \begin{bmatrix} \frac{\partial \varepsilon_t}{\partial \beta_1} \\ \frac{\partial \varepsilon_t}{\partial \beta_2} \\ \frac{\partial \varepsilon_t}{\partial \gamma} \end{bmatrix} = \begin{bmatrix} x_{1t} \\ \frac{1}{2}(x_{2t} - \gamma)^{-2} \\ -\beta_2(x_{2t} - \gamma)^{-3} \end{bmatrix}.$$

Solving  $\sum z_t \varepsilon_t = 0$  yields ML estimates of  $(\beta_1, \beta_2, \gamma)$ .  $\hat{\sigma}^2 = \frac{1}{T} \sum \hat{\varepsilon}_t^2$ .

$H_0 : \gamma = 0$ . Under  $H_0$

$$\begin{aligned} y_t &= \beta_1 x_{1t} + \beta_2 \frac{1}{2x_{2t}^2} + \varepsilon_t \\ \varepsilon_t &= y_t - \beta_1 x_{1t} - \frac{\beta_2}{2x_{2t}^2}. \end{aligned}$$

ML estimation of restricted model is least squares regression of  $y_t$  on  $x_{1t}$ ,  $x_{2t}^{-2}$ .

Then calculate

$$\hat{\varepsilon}_t = y_t - \hat{\beta}_1 x_{1t} - \hat{\beta}_2 \frac{1}{2x_{2t}^2}.$$

When  $\gamma = 0$ ,  $z_t = \begin{bmatrix} x_{1t} \\ \frac{1}{2x_{2t}^2} \\ -\frac{\beta_2}{x_{2t}^3} \end{bmatrix}$ . LM test requires regressing  $\hat{\varepsilon}_t$  on  $z_t$ , calculate  $TR^2$ . Reject  $H_0$  if  $TR^2 > \chi_\alpha^2(1)$ .

LM easier than LR as you do not have to do nonlinear least squares maximization.

## 2. Distributed lags and transformations

1. Rewrite the distributed lag

$$B(L)x_t = \beta_0 x_t + \beta_1 x_{t-1} + \dots + \beta_m x_{t-m}$$

in the form

$$\delta_0 x_t + \delta_1 \Delta x_t + \delta_2 \Delta x_{t-1} + \dots + \delta_m \Delta x_{t-m+1}$$

1. Show that  $\delta_0 = B(1)$ , the total effect. What is the relationship between the  $\delta$ 's and the  $\beta$ 's?

**Answer.**

$$\beta_0 x_t + \beta_1 x_{t-1} + \dots + \beta_m x_{t-m} =$$

$$\delta_0 x_t + \delta_1 (x_t - x_{t-1}) + \delta_2 (x_{t-1} - x_{t-2}) + \dots + \delta_m (x_{t-m+1} - x_{t-m})$$

$$\therefore \beta_m = -\delta_m$$

$$\delta_m = -\beta_m$$

$$\beta_{m-1} = \delta_m - \delta_{m-1}$$

$$\delta_{m-1} = -(\beta_{m-1} + \beta_m)$$

$$\beta_{m-2} = \delta_{m-1} - \delta_{m-2}$$

$$\delta_{m-2} = -(\beta_{m-2} + \beta_{m-1} + \beta_m)$$

$$\vdots$$

$$\beta_1 = \delta_2 - \delta_1$$

$$\delta_1 = -(\beta_1 + \dots + \beta_m)$$

$$\beta_0 = \delta_0 + \delta_1$$

$$\delta_0 = \beta_0 + \beta_1 + \dots + \beta_m = B(1) \text{ total effect.}$$

2. If you write the second form as

$$\delta_0^+ x_{t-1} + \delta_1^+ \Delta x_t + \delta_2^+ \Delta x_{t-1} + \dots + \delta_m^+ \Delta x_{t-m+1}$$

how are the  $\delta^+$ 's related to the  $\beta$ 's?

**Answer.**

$$\delta_0^+ x_{t-1} + \delta_1^+ (x_t - x_{t-1}) + \delta_2^+ (x_{t-1} - x_{t-2}) + \dots + \delta_m^+ (x_{t-m+1} - x_{t-m})$$

$$\beta_m = -\delta_m^+ \quad \therefore \delta_m^+ = -\beta_m$$

$\vdots$

$$\beta_2 = \delta_3^+ - \delta_2^+ \quad \therefore \delta_2^+ = -(\beta_2 + \dots + \beta_m)$$

$$\beta_0 = \delta_1^+ \quad \therefore \delta_1^+ = \beta_0$$

$$\beta_1 = \delta_0^+ - \delta_1^+ + \delta_2^+ \quad \therefore \delta_0^+ = \beta_1 + \beta_0 + \beta_2 + \dots + \beta_m = B(1)$$

*Note: coeff. on levels variable unchanged (still total effect).*

2. Consider the Autoregressive Distributed Lag Model

$$A(L)y_t = \lambda + B(L)x_t + \varepsilon_t$$

where

$$A(L) = 1 - \alpha_1 L - \dots - \alpha_m L^m$$

$$B(L) = \beta_0 + \beta_1 L + \dots + \beta_n L^n.$$

1. Show that you can rewrite this model in the error correction form

$$\Delta y_t = \lambda + \gamma_0 y_{t-1} + \sum_{i=1}^{m-1} \gamma_i \Delta y_{t-i} + \delta_0 x_{t-1} + \sum_{i=1}^n \delta_i \Delta x_{t-i+1} + \varepsilon_t.$$

What is the interpretation of  $\gamma_0$  and  $\delta_0$  and hence of  $\delta_0/\gamma_0$ ? [Hint: use the results you derived in the first question of this section]

**Answer.**

$$\begin{aligned} (y_t - y_{t-1}) &= \lambda + \gamma_0 y_{t-1} + \sum_{i=1}^{m-1} \gamma_i (y_{t-i} - y_{t-i-1}) \\ &\quad + \delta_0 x_{t-1} + \sum_{i=1}^n \delta_i (x_{t-i+1} - x_{t-i}) + \varepsilon_t \end{aligned}$$

$$y_t = \lambda + (\gamma_0 + 1)y_{t-1} + \sum_{i=1}^{m-1} \gamma_i (y_{t-i} - y_{t-i-1}) + B(L)x_t + \varepsilon_t$$

where  $\delta_0 = B(1)$  and  $\delta_1 = \beta_0$ ,  $\delta_i = -\sum_{j=i}^n \beta_j$  from results above.

$$y_t - (1 + \gamma_0)y_{t-1} - \sum_{i=1}^{m-1} \gamma_i(y_{t-i} - y_{t-i-1}) = \lambda + B(L)x_t + \varepsilon_t.$$

$$LHS = y_t - (1 + \gamma_0)y_{t-1} - \gamma_1(y_{t-1} - y_{t-2}) - \gamma_2(y_{t-2} - y_{t-3}) - \dots - \gamma_{m-1}(y_{t-m+1} - y_{t-m})$$

$$\begin{aligned} \therefore \quad & -\alpha_m = \gamma_{m-1} & \gamma_{m-1} &= -\alpha_m \\ & -\alpha_{m-1} = \gamma_{m-2} - \gamma_{m-1} & \gamma_{m-2} &= -(\alpha_{m-1} + \alpha_m) \\ & -\alpha_2 = \gamma_1 - \gamma_2 & \gamma_1 &= -(\alpha_2 + \dots + \alpha_m) \\ & -\alpha_1 = -(1 + \gamma_0) - \gamma_1 & \gamma_0 &= -(1 - \alpha_1 - \alpha_2 - \dots - \alpha_m) \\ & & &= -A(1) \end{aligned}$$

$$\frac{\delta_0}{\gamma_0} = -\frac{B(1)}{A(1)}, \text{ this is minus the steady state effect of a change in } x_t \text{ on } y_t.$$

2. If you estimate the original model and the model in error correction form, what is the relationship between the two sets of estimates and the two sets of residuals?

**Answer.** *As these are non-singular linear transforms, the estimates have the same relationships as above and the residuals are the same.*