

Discussion of:

The Implied Equity Term Structure

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The bigger picture

Absent arbitrage opportunities, it exists an SDF (M) s.t. the price (P_t) of an asset that delivers the cashflows $\{D_{t+i}\}_{i=1}^{\infty}$ can be decomposed as

$$P_t = \underbrace{\sum_{i=1}^T \mathbb{E}_t [M_{t,t+i} D_{t+i}]}_{\text{price of ST asset} =: P_t^{(1:T)}} + \underbrace{\sum_{i=T+1}^{\infty} \mathbb{E}_t [M_{t,t+i} D_{t+i}]}_{\text{price of LT asset} =: P_t^{(T+1:\infty)}} = \sum_{i=1}^{\infty} \underbrace{P_t^{(i)}}_{\text{price of "bullet" CF} \equiv P_t^{(ii)}}$$

The **dividend strip prices** $P_t^{(i)}$ are very salient since economic theory has stringent predictions on the shape of the term structure of

$$\mathbb{E} \left[R_{t,t+i}^{(i)} \right] := \mathbb{E} \left[\frac{D_{t+i}}{P_t^{(i)}} \right] \quad \text{and} \quad \mathbb{E} \left[R_{t,t+1}^{(i)} \right] := \mathbb{E} \left[\frac{P_{t+1}^{(i-1)}}{P_t^{(i)}} \right]$$

E.g. monotonically increasing risk premia in Habit models, increasing but with horizontal asymptote in vanilla LRR, flat in disaster models.

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The bigger picture cont'd

Problem: $P_t^{(i)}$ s are NOT generally available.

Solutions:

1. Use implied $P_t^{(i)}$ from derivative markets (index options, index futures, dividend futures): e.g. van Binsbergen-Kojen 2017, van Binsbergen-Brandt-Kojen 2012, Bansal-Yaron-Miller-Song 2021
2. Affine modelling of SDF, price, and dividend processes: Kelly-Giglio-Kozak 2021
3. **This paper:** estimates directly $\mathbb{E} [R_{t,t+i}^{(i)}]$ using the fact that

$$P_t = \sum_{i=1}^{\infty} \mathbb{E}_t [M_{t,t+i} D_{t+i}] \equiv \sum_{i=1}^{\infty} \frac{\mathbb{E}_t [D_{t+i}]}{\mathbb{E}_t [R_{t,t+i}^{(i)}]}$$

⇒ clever and fresh idea - I like it!

But needs:

- i. forecasting model/method for $\mathbb{E}_t [D_{t+i}]$ (analysts' forecasts + AR(1))
- ii. Enough restrictions for identification of $\left\{ \mathbb{E}_t [R_{t,t+i}^{(i)}] \right\}_{i=1}^{\infty}$ (truncation at $T+$ parametric form)

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Comments

I. Restrictions on $\mathbb{E}_t \left[R_{t,t+i}^{(i)} \right]$: time & cross-sectional variation + monotonicity

- Define (as in the paper): $r_{t,i} \equiv \mathbb{E}_t \left[R_{t,t+i}^{(i)} \right]^{\frac{1}{i}} - 1$ (the annualized discount rate)
- The authors assume the parametric form ("kind of" Nelson-Siegel):

$$r_{t,i} - r_{t,i}^f = \beta_1 - \beta_2 \left[\frac{1 - \exp(-i/\lambda)}{i/\lambda} \right] \quad \forall i \leq T, \quad r_{t,i} = r_{t,T} \quad \forall i > T; \quad \forall t$$

and estimate β 's but calibrate λ (☺)

1. Conditional risk premia are constant (cross-sectionally and in time)

... but they don't need to be for identification!

- If parameters are fixed in time, there are enough d.o.f. to allow them to vary across assets (3 periods are enough).
- If parameters are fixed cross-sectionally, there are enough d.o.f. to allow them to vary across time (3 assets are enough)

Note: the authors show that they vary both in time (BC) and across characteristics

⇒ I'd model variation in β 's and λ formally: e.g. i) linear functions of characteristics (to capture cross-sectional variation) and aggregated state variables (to capture time variation) or ii) (Bayesian) latent TVPs

- more than 500 assets and 40 years of data make it highly feasible!

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2. Conditional risk premia are strictly monotonic

- Why not full Nelson-Siegel?
- Only one more parameter, and curvature likely to depend on characteristics too.

⇒ I'd like formal estimation and selection of the β 's and λ .

- hard to take the slope findings at face value without **formal testing** of e.g. $H_0 : \beta_2 \neq 0$

Note: the estimation problem can be written as a NLS (just divide P_t by D_t and add an error)

⇒ can use standard methods (GMM, MLE) for inference. Can also test formally which characteristics drive the heterogeneity.

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II. Restrictions on $\mathbb{E}_t [D_{t+i}]$

Cash-flows are modelled observing that:

$$D_t = \underbrace{BE_{t-1}}_{\text{book equity}} \times \left(ROE_t - \underbrace{\%BE_t}_{\text{BE growth}} \right)$$

And assuming:

1. $\mathbb{E}_t [D_{t+i}] = \text{analysts' forecast } \forall i < 6 \text{ years}$
2. AR(1) for both ROE and BE (stationarity? maybe typo: $\%BE$ instead?) $\forall i > 6 \text{ years}$ with coefficients from a pooled regression.

⇒ Same identical expected growth from 6 years onward... but don't we expect e.g. "growth" and "value" stocks to have different growth rates?

Issue: If growth rates are different, the above will generate spurious heterogeneity in discount rates.

⇒ I'd estimate firm/portfolio specific ARMA processes (again, tons of dof)

Note: $\%BE_t$ proxied by sales growth – why? BE is "observable"

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III. Inference and interpretation

1. Many of the term structure estimates have discount rates in the $[-8\%, +8\%]$ range, and rates tend to be negative for maturities below ≈ 6 years...

⇒ needs proper confidence bands to trust the conclusions on the slope (again, cast the estimation as a big GMM – including the ARs – and use e.g. the Delta method)

Note: the negative discount rates corresponding to analyst forecasts suggest the analysts' medium run forecast are too low relative / observed prices too high – very different story...

⇒ Large literature on biases in analyst forecasts' (cf. Bradshaw-Ertimur-O'Brien 2016)

2. The estimated $\mathbb{E}_t \left[R_{t,t+i}^{(i)} \right]^{\frac{1}{i}} \forall i$ is nothing but a (filtered) asset pricing model

⇒ **needs to show that it is a good pricing model** if we have to believe the estimates. Report the pricing errors and standard fit metrics (MAPE, R^2 , J -stat etc.)

Note: might be hard to do well without heterogeneity – all assets have a $\beta^{(i)} = 1$

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⇒ it's ok, but the direct comparison should be careful (e.g. Kelly-Giglio-Kozak (2021))

Note: from the literature we know what to expect for the latter quantities in canonical models but not the former ⇒ simulate LRR and Habit models to give us a benchmark.

4. In the absence of arbitrage opportunities:

$$\underbrace{\mathbb{E} \left[\frac{P_t^{(i)}}{P_t} \right]}_{\text{values share of strip}} \equiv \underbrace{\text{Cov} \left(M_{t,t+i}; \frac{D_{t+i}}{P_t} \right)}_{\text{risk adjustment}} + \underbrace{\mathbb{E} [M_{t,t+i}] \mathbb{E} \left[\frac{D_{t+i}}{P_t} \right]}_{\text{risk neutral valuation}}$$

⇒ Use a sound SDF (e.g. Kozak-Nagel-Santosh 2020, Huang-Bryzgalova-Julliard 2022) to verify your findings.

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Note: from the literature we know what to expect for the latter quantities in canonical models but not the former ⇒ simulate LRR and Habit models to give us a benchmark.

4. In the absence of arbitrage opportunities:

$$\underbrace{\mathbb{E} \left[\frac{P_t^{(i)}}{P_t} \right]}_{\text{values share of strip}} \equiv \underbrace{\text{Cov} \left(M_{t,t+i}; \frac{D_{t+i}}{P_t} \right)}_{\text{risk adjustment}} + \underbrace{\mathbb{E} [M_{t,t+i}] \mathbb{E} \left[\frac{D_{t+i}}{P_t} \right]}_{\text{risk neutral valuation}}$$

⇒ Use a sound SDF (e.g. Kozak-Nagel-Santosh 2020, Huang-Bryzgalova-Julliard 2022) to verify your findings.

Conclusion

Conclusion

- (+) A very good and novel idea, and I really enjoyed reading the paper.
- (+) Results are (potentially) very salient and interesting.
- (-) But I would like the analysis to be much more econometrically formal to fully trust the findings (relatively “easy” to fix).