

PS7

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Binary response model

Random effects probit (xtprobit)

Probit case: Assume $y_{it}^* = x'_{it}\beta + \varepsilon_{it}$

Assuming $\varepsilon_{it} \sim \mathcal{N}(0, 1)$ i.i.d we get the usual probit model that can be estimated by MLE.

If we have a one factor error, ε_{it} , we have: $\varepsilon_{it} = \alpha_i + \nu_{it}$

To find a MLE, we need to make some extra assumptions on the disturbances, that will give us the **random effects probit model**:

1. $y_{it}^* = x'_{it}\beta + \varepsilon_{it} = x'_{it}\beta + \alpha_i + \nu_{it}$ with $y_{it} = 1_{y_{it}^* > 0}$.
2. $\alpha_i | X \sim \mathcal{N}(0, \sigma_\alpha^2)$ i.i.d over i .
3. $\nu_{it} | X, \alpha_i \sim \mathcal{N}(0, 1)$ i.i.d over i and t .

Rk. 1 Assumptions 2 and 3 imply that the correlation between two disturbances of the same individual is constant over time. This is sometimes called an equi-correlation assumption. Assumption 2 is the **random effects** assumption. It states that the distribution of α_i conditional on X does not depend on X .

Rk. 2 Here the panel data is balanced. So we can write: $E(\varepsilon\varepsilon' | X) = I_n \otimes \Omega$

where $\Omega_{T \times T} = (1 + \sigma_\alpha^2) \cdot \begin{pmatrix} 1 & \rho & \dots \\ & \ddots & \\ \dots & \rho & 1 \end{pmatrix}$, with $\rho = \frac{\sigma_\alpha^2}{1 + \sigma_\alpha^2}$.

Then, $P(y_{it} = 1 | X, \alpha_i) = \Phi(x'_{it}\beta + \alpha_i)$ by A1 and A3.

Thus, $P(y_{it} = 1 | x_{it}) = \int \Phi(x'_{it}\beta + \alpha) f_{\alpha|x_{it}}(\alpha) d\alpha$.

Or, $P(y_{it} = 1 | x_{it}) = \int \Phi(x'_{it}\beta + \alpha) f_\alpha(\alpha) d\alpha$ by A2.

Thus $P(y_{it} = 1 | x_{it}) = \int \Phi(x'_{it}\beta + \alpha) \frac{1}{\sigma_\alpha} \varphi\left(\frac{\alpha}{\sigma_\alpha}\right) d\alpha$ by A2.

However, in general, $P(y_{i1}, \dots, y_{iT} | X_i) \neq \prod_t P(y_{it} | x_{it})$

As conditional on X_i , the random part of the y_{it} s are the $\varepsilon_{it} = \alpha_i + \nu_{it}$ which share the same α_i .

Hence we have to apply the same argument to the likelihood of all the observations of an individual, i :

$P(y_{i1}, \dots, y_{iT} | x_i, \alpha_i) = P(y_i | x_i, \alpha_i) = \prod_t P(y_{it} | x_i, \alpha_i) = \prod_t P(y_{it} | x_{it}, \alpha_i)$ by A3.

So, $P(y_i | x_i) = \int \prod_t P(y_{it} | x_i, \alpha) f_{\alpha|x_i}(\alpha) d\alpha = \int \prod_t P(y_{it} | x_i, \alpha) f_\alpha(\alpha) d\alpha$ by

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A2.

And, $P(y_i|x_i) = \int \prod_t P(y_{it}|x_{it}, \alpha) \frac{1}{\sigma_\alpha} \varphi\left(\frac{\alpha}{\sigma_\alpha}\right) d\alpha$ by A2.

Note that by A3, we have: $P(y_{it}|x_{it}, \alpha) = \Phi(x'_{it}\beta + \alpha)^{y_{it}} (1 - \Phi(x'_{it}\beta + \alpha))^{1-y_{it}}$

Thus, we obtain the likelihood for the full sample:

$L(y_1, \dots, y_n|x_1, \dots, x_n) = \prod_i P(y_i|x_i)$ by A2 and A3.

$L(y_1, \dots, y_n|x_1, \dots, x_n) = \prod_i \int \prod_t P(y_{it}|x_{it}, \alpha) \frac{1}{\sigma_\alpha} \varphi\left(\frac{\alpha}{\sigma_\alpha}\right) d\alpha$

Under the assumptions of the random effects probit (one factor model), this is a parametric and fully efficient estimator. We obtain estimators of β and σ_α . The estimator $\hat{\sigma}_\alpha$ can be used to test the presence of individual specific effects, α_i , under the random effects assumption.

Extension to ARMA errors

Now we assume:

1. $y_{it}^* = x'_{it}\beta + \varepsilon_{it} = x'_{it}\beta + \alpha_i + \nu_{it}$ with $y_{it} = 1_{y_{it}^* > 0}$.
2. $\alpha_i|X \sim \mathcal{N}(0, \sigma_\alpha^2)$ i.i.d over i .
3. $\nu_{it}|X, \alpha_i \sim ARMA(p, q)$.

$\nu_{it}|X, \alpha_i \sim ARMA(p, q)$ means that:

$\nu_{it} = \sum_p \alpha_p \nu_{it-p} + \sum_q \gamma_q \zeta_{it-q}$ where ζ_{it} are iid over i and t .

We get a general likelihood:

$$L = \prod_{i=1}^n \underbrace{\left(\int \dots \int_{a_{it}}^{b_{it}} \dots \int f_{\varepsilon_{T_i}}(\varepsilon_{i1}, \dots, \varepsilon_{iT_i}) \cdot d\varepsilon_{T_i} \right)}_{T_i \text{ integrals}}$$

where if $y_{it} = 1$ then $a_{it} = -x'_{it}\beta$ and $b_{it} = +\infty$.

while if $y_{it} = 0$ then $b_{it} = -x'_{it}\beta$ and $a_{it} = -\infty$.

The integral does not simplify in this case because of the temporal relationship between the transitory shocks ν_{it} . We have to use simulation based methods.

We can use SML, simulated maximum likelihood:

1. The parameters of interest are $\theta_{(K+p+q+1) \times 1} = (\beta', \nu', \gamma', \sigma_\alpha)'$.
2. We maximize $\tilde{L}(\theta, R) = 1/n \sum_i^n \ln(\tilde{l}(y_i, \theta, R))$ where R is the number of simulations and \tilde{l} the simulated likelihood.
3. We first draw $R, T \times 1$ uniform random vectors \tilde{u}_r . We will keep fixed these vectors in all the following process.
4. Let $\theta^{(k)}$ be the value of the parameters at iteration k . To obtain \tilde{L} we proceed as follows:

- From our assumptions about the pdf of ε and our parameters' values, $\theta^{(k)}$, we get the simulated disturbances: $\tilde{\varepsilon}_{i,r}^{(n)} = F_{\varepsilon_i}^{-1}(\tilde{u}_r, \theta^{(n)})$. This gives us: $\tilde{y}_{i,r}^{*(n)}$ and $\tilde{y}_{i,r}^{(n)}$.
 - Using the R simulations we get: $\tilde{l}(y_i, \theta^{(n)}, R)$ for each i and we compute $\tilde{L}(\theta^{(n)}, R)$.
5. We iterate step 4 among the possible values of θ to maximize the simulated log-likelihood.

Random effects tobit

We assume:

$$y_{it}^* = x_{it}'\beta + \varepsilon_{it} = x_{it}'\beta + \alpha_i + \nu_{it}$$

$$y_{it} = 1_{y_{it}^* > 0} * y_{it}^*$$

By the same arguments as before we have:

$$f(y_i|X_i) = \int f(y_i|X_i, \alpha) \cdot f_\alpha(\alpha|X_i) d\alpha$$

Assuming $\alpha_i|X \sim \mathcal{N}(0, \sigma_\alpha^2)$ i.i.d over i , this becomes:

$$f(y_i|X_i) = \int f(y_i|X_i, \alpha) \cdot \frac{1}{\sigma_\alpha} \varphi\left(\frac{\alpha}{\sigma_\alpha}\right) d\alpha$$

Assuming $\nu_{it}|X, \alpha_i \sim \mathcal{N}(0, \sigma_\nu^2)$ i.i.d over i and t , we get:

$$f(y_i|X_i) = \int \prod_t f(y_{it}|x_{it}, \alpha) \cdot \frac{1}{\sigma_\alpha} \varphi\left(\frac{\alpha}{\sigma_\alpha}\right) d\alpha$$

- If $y_{it} = 0$ then $f(y_{it}|x_{it}, \alpha) = \Phi\left(\frac{-x_{it}'\beta - \alpha}{\sigma_\nu}\right)$
- If $y_{it} > 0$ then $f(y_{it}|x_{it}, \alpha) = \frac{1}{\sigma_\nu} \varphi\left(\frac{y_{it} - x_{it}'\beta - \alpha}{\sigma_\nu}\right)$

By independence of the observations over the individuals i we get the likelihood as:

$$L = \prod_{i=1}^n f(y_i|X_i) = \prod_{i=1}^n \left[\int \prod_t f(y_{it}|x_{it}, \alpha) \cdot \frac{1}{\sigma_\alpha} \varphi\left(\frac{\alpha}{\sigma_\alpha}\right) d\alpha \right].$$