

EC402 (2023/24)

RAGVIK'S SPEAKING NOTES FOR CLASS

Vassilis PS 9 (WK 10, AT)

INTUITION FOR IV

Say $y_t = x_t' \beta + \varepsilon_t$ and $E[x_t \varepsilon_t] \neq 0$ for $t=1, \dots, T$.

Find a consistent estimator for β .

QUESTION 1

$$(Q1) \quad y = X_A \beta_A + X_B \beta_B + \varepsilon$$

where X_A is $T \times (K-1)$ and

X_B is $T \times 1$ and $E[X_B' \varepsilon] \neq 0$.

We have $A1, A2, A3 \in \mathcal{M}_i, X_A, A4 \in \mathcal{M}_2, A5 \in \mathcal{N}$.

$$(a) \quad \hat{\beta}_{OLS} := (X'X)^{-1} X'y \text{ where } X := [X_A \ X_B].$$

$$\hat{\beta}_{IGLS} := (X' \tilde{\Omega} X)^{-1} X' \tilde{\Omega} y.$$

$$\text{let } \hat{\beta} := (\hat{\beta}_A, \hat{\beta}_B)', \hat{\beta}_{OLS} := \begin{pmatrix} \hat{\beta}_{A, OLS} \\ \hat{\beta}_{B, OLS} \end{pmatrix} \text{ and} \\ \hat{\beta}_{IGLS} := \begin{pmatrix} \hat{\beta}_{A, IGLS} \\ \hat{\beta}_{B, IGLS} \end{pmatrix}.$$

"Explain carefully... unbiasedness"

Consider $(\hat{\beta}_{OLS} - \beta_A) = (X_A' M_B X_A)^{-1} X_A' M_B \varepsilon$ where $M_B := [I_T - X_B (X_B' X_B)^{-1} X_B']$.

Now, $E[\hat{\beta}_{OLS} - \beta_A] \stackrel{LIE}{=} E[(X_A' M_B X_A)^{-1} X_A' M_B E[\varepsilon | X]] \stackrel{A3.X_B}{\neq} 0$, even under $A3.Rmi.X_A$

Similarly,

$(\hat{\beta}_{IGLS} - \beta_A) = (\tilde{X}_A' \tilde{M}_B \tilde{X}_A)^{-1} \tilde{X}_A' \tilde{M}_B \tilde{\varepsilon}$ where

$$\tilde{X}_A := \tilde{\Sigma}^{1/2} X_A$$

$$\tilde{\varepsilon} := \tilde{\Sigma}^{1/2} \varepsilon$$

$$\tilde{M}_B := [I_T - \tilde{\Sigma}^{1/2} X_B ((\tilde{\Sigma}^{1/2} X_B)' (\tilde{\Sigma}^{1/2} X_B))^{-1} (\tilde{\Sigma}^{1/2} X_B)']$$

So $E[\hat{\beta}_{IGLS} - \beta_A] \stackrel{LIE}{=} E[(\tilde{X}_A' \tilde{M}_B \tilde{X}_A)^{-1} \tilde{X}_A' \tilde{M}_B \tilde{\Sigma}^{1/2} E[\varepsilon | X]] \stackrel{A3.X_B}{\neq} 0$
even under $A3.Rmi.X_A$.

"Explain carefully... Consistency."

For some $K-1$ by $K-1$ finite P.D. matrix J ,

$$\begin{aligned} \text{plim}_{T \rightarrow \infty} (\hat{\beta}_{OLS}^A - \beta_A)^{ST} &= \left[\text{plim}_{T \rightarrow \infty} (\bar{T}' X_A' M_B X_A) \right]^{-1} \text{plim}_{T \rightarrow \infty} \bar{T}' X_A' M_B \varepsilon \\ &= J \cdot \text{plim}_{T \rightarrow \infty} \bar{T}' \left[X_A' \varepsilon - X_A' X_B (X_B' X_B)^{-1} X_B' \varepsilon \right] \end{aligned}$$

$\neq 0$ even if $\text{plim}_{T \rightarrow \infty} (\bar{T}' X_A' \varepsilon) \stackrel{A3 Rmi. X_A}{=} 0$ since

$\text{plim}_{T \rightarrow \infty} (\bar{T}' X_B' \varepsilon) \stackrel{A3. X_B}{\neq} 0$ unless $\text{plim}_{T \rightarrow \infty} \frac{1}{T} (X_A' X_B) = 0$.

Similarly, since $\text{plim}_{T \rightarrow \infty} \bar{T}' (\bar{\Sigma}^{1/2} X_B)' \bar{\Sigma}^{1/2} \varepsilon \neq 0$, we also have that $\text{plim}_{T \rightarrow \infty} (\hat{\beta}_{OLS}^A - \beta_A) \neq 0$.

ASK YOURSELF: What's the key message of these slides?

(b) Consider $\hat{\beta}_{IV} := (W'X)^{-1}W'y$ where $A1, A2, A3 Rmi. X_A, A4 GM$
 W is a $T \times K$ matrix of instruments s.t. $\text{rank}(W'X) = K$

and (i) $\lim_{T \rightarrow \infty} \bar{T}' W' \varepsilon = 0$; and

(ii) $\lim_{T \rightarrow \infty} \bar{T}' W' X = \Sigma_{WX}$, where Σ_{WX} is a finite non-singular $K \times K$ matrix.

[Which is validity and which is relevance?]

In the over-identified case (such as if $K_Z > 1$),

$W'X$ is $[K_Z + (K-1)] > K$ by K matrix.

[So what can we do? Throw instruments away? Nope...]

Consider the following scheme:

1. Define $Q := [X_A \ Z]$, a T by $(K-1) + K_Z$ matrix where $\text{rank}(Q'X) = K$ and

$$(i) \lim_{T \rightarrow \infty} T^{-1} Q' \varepsilon = 0$$

$$(ii) \lim_{T \rightarrow \infty} T^{-1} Q' X = \sum_{QX}, \text{ a full column rank matrix.}$$

2. Choose $W := Q(Q'Q)^{-1}Q'X$

3. Obtain $\hat{\beta} := (W'W)^{-1}W'y$ [careful of "auto" $SE \hat{\beta}$]

$$\begin{aligned} &= [X'Q(Q'Q)^{-1}Q'Q(Q'Q)^{-1}Q'X]^{-1} X'Q(Q'Q)^{-1}Q'y \\ \equiv \hat{\beta}_{GIV} &:= [X'Q(Q'Q)^{-1}Q'X]^{-1} X'Q(Q'Q)^{-1}Q'y \end{aligned}$$

$$\equiv \hat{\beta}_{IV} := (W'X)^{-1}W'y.$$

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- Asymptotics (as $T \rightarrow \infty$) :

$$\sqrt{T} (\hat{\beta}_{IV} - \beta) \xrightarrow{d} \text{MVN} \left(0, \sigma^2 \left[\Sigma_{XW} \Sigma_{WW}^{-1} \Sigma_{WX} \right] \right) \text{ as } T \rightarrow \infty$$

$$\text{where } \Sigma_{XW} := \text{plim}_{T \rightarrow \infty} \frac{1}{T} X'W,$$

$$\Sigma_{WW} := \text{plim}_{T \rightarrow \infty} \frac{1}{T} W'W, \text{ and}$$

$$\Sigma_{WX} := \Sigma_{XW}' \text{ are all finite and invertible matrices.}$$

(Also, σ^2 is as defined in the question.)

- Estimation of variance :

$$\hat{\text{var}}(\hat{\beta}_{IV}) := \hat{\sigma}^2 \left[(X'W) (W'W)^{-1} (W'X) \right]$$

$$\text{where } \hat{\sigma}^2 := \frac{1}{T-k} (y - X \hat{\beta}_{IV})' (y - X \hat{\beta}_{IV})$$