Structural Instabilities in Factor Models for Time Series

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Outline	2		

- Co-movements of (large) number of observed time series modelled in terms of small number of unobserved common factors.
- Literature on high-dimensional GDFMs focusses on case where parameters are stable; in particular, where factor loadings are time-invariant.

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Outline of today's discussion

Debate: do structural instabilities matter?

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- Model: with/without structural breaks
- Roadmap: what the key players are saying

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- Debate: do structural instabilities matter?
- Model: with/without structural breaks
- Roadmap: what the key players are saying
- Extension: what I plan to do

I. Debate			

I. What is the debate about?

Image: A mathematical states and a mathem

Do structural instabilities matter?

NAY!

- Stock & Watson (JASA 2002)
- Bates, Plagborg-Moller, Stock & Watson (JoE 2013)

Claim:

Can consistently estimate the factor space despite structural instabilities

... provided certain conditions hold.

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Extended model

IV. Bates et al. (2013)

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Aye!

- Breitung & Eickmeier (JoE 2011)
- Corradi & Swanson (Forthcoming, JoE)
- Chen, Dolado & Gonzalo (JoE 2014)
- Han & Inoue ('metric Theory 2014)
- Baltagi, Kao & Wang (Manuscript, 2015)
- Yamamoto & Tanaka (JoE 2015)

Claim:

Stock & Watson conditions not realistic.

II. Standard model		

II. The standard model, with no breaks

Image: A mathematical states and a mathem

Standard model: $X_t = \chi_t + e_t = \Lambda_t F_t + e_t$

For t = 1, ..., T

- Observed series: X_t ; $dim(X_t) = N$
- Common component: χ_t ; $dim(\chi_t) = N$
- ▶ Idiosyncratic error: e_t ; $dim(e_t) = N$, with Σ_e not necessarily diagonal
- Common factors: F_t ; $dim(F_t) = r$, where $r \ll N$
- ► Factor loadings: Λ_t ; an $(N \times r)$ matrix, with Λ_0 fixed

Notation alert:

 $\left\{ \begin{array}{l} "N": total number of observed variables (sometimes denoted by "p") \\ "T": total number of time periods (sometimes denoted by "n") \end{array} \right.$

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Question:

How can we estimate this model? Can we use PCA?

Can use PCA under certain conditions

Under certain assumptions (to be discussed), as $N \rightarrow \infty$, the following are equivalent:

- ▶ PCA projection of X_t onto space spanned by the *r* first principal components;
- Factor Analysis projection of X_t onto space spanned by the r factors F_t .

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Emphasis:

- Critical that $N \rightarrow \infty$ for equivalence between PCA and Factor Analysis to hold.
- Critical that certain assumptions are not violated. We discuss these now...

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Main assumptions; Bai & Ng (Econometrica 2002)

- 1. Pervasiveness of the factors:
 - $\Lambda'_0 \Lambda_0 / N \rightarrow D$ as $N \rightarrow \infty$ for some positive definite $(r \times r)$ matrix D.
- Factors affect (most or) all of the observed series, even when the number of series grows very large.

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- Factors affect (most or) all of the observed series, even when the number of series grows very large.
- *F_t* can be assumed stationary for simplicity.
- It is equivalent to assume that:
 - Σ_{χ} has all *r* eigenvalues diverging as $N \to \infty$.

- 2. Mild cross-sectional correlation in idiosyncratic component:
 - $\frac{1}{N}\sum_{i,j=1}^{N} |Cov(e_{i,t}, e_{j,t})| \le M$ for some $M \in (0, \infty)$.
- ▶ This assumption allows for the generalised/approximate factor structure; yet...

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- it prevents combinations of idiosyncratic terms from having unusually large variation; otherwise it would be difficult to identify genuine factors.
- We can assume that e_t is stationary. No restriction on temporal dependence.
- Heteroscedasticity in both dimensions is also allowed.
- It is (almost) equivalent to assume that:
 - Σ_e has all eigenvalues bounded for any N.

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The assumptions on eigenvalues of Σ_{χ} and Σ_{e} imply...

- ▶ ...that the r^{th} largest eigenvalue of Σ_x diverges as $N \to \infty$; and
- the $(r+1)^{th}$ largest eigenvalue of Σ_x stays bounded for any N.

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In other words:

PCA may be used, alongside any number of existing criteria for estimating r, in order to consistently recover the space spanned by the r factors.

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In other words:

PCA may be used, alongside any number of existing criteria for estimating r, in order to consistently recover the space spanned by the r factors.

As
$$N, T \to \infty$$

 $\|\hat{F}_t - H'F_t\| = O_p\left(\max\left\{\frac{1}{\sqrt{N}}, \frac{1}{\sqrt{T}}\right\}\right)$

The rate of convergence is determined by the smaller of N or T, and thus depends on the panel structure.

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	III. Extended model		

III. The extended model, with breaks in factor loadings



A FM with breaks in the loadings can be re-parameterized as another FM with constant loadings but a larger set of factors.

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- Consider, for instance, a one-factor model with a structural break:

$$X_{it} = egin{cases} \lambda_{1i}f_t + e_{it}, t \leq \kappa \ \lambda_{2i}f_t + e_{it}, t > \kappa \end{cases}$$



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• Define
$$g_{1t} = \begin{cases} f_t, t \leq \kappa \\ 0, t > \kappa \end{cases}$$
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• Then, we have $X_{it} = \lambda_{1i}g_{1t} + \lambda_{2i}g_{2t} + e_{it}$, an equivalent stable model.

Image: A mathematical states and a mathem

Graphical illustration with one-factor model





Existing criteria to determine the number of factors, r, will fail.

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Bai & Ng (Econometrica 2002)
Onatski (RES 2010)
Alessi, Barigozzi & Capasso (Stat. & Prob. Letters 2010)
Ahn & Horenstein (Econometrica 2013)
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Existing methods to estimate the space spanned by the factors will fail.

∫PCA | Kalman Filter

Implications:

Pseudo-factors - not necessarily interpretable
Forecasting - performance may deteriorate
Structural analysis - "impulse responses" not recoverable

	IV. Bates et al. (2013)	

IV. (Moderate) instabilities do not matter

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Model with instability:

$$X_t = \Lambda_t F_t + e_t$$
$$\Lambda_t = \Lambda_0 + \xi_t$$

• ξ_t : a random process of dimension $N \times r$

For example, a single abrupt break at a common date κ : Specifically, let r = 1, and $\Delta \in \mathbb{R}^N$ be a shift parameter.

We define:

$$\xi_t = \begin{cases} 0 \text{ for } t = 1, ..., \kappa \\ \Delta \text{ for } t = \kappa + 1, ..., T \end{cases}$$

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Question:

- Under what (extra) conditions is PCA estimator still consistent...
- ...with the same convergence rate as for stable model?

What additional assumptions do we need?

Bates et al. (JoE 2013):

- Factor loading innovations (breaks)
 - 1. Limited dependence between breaks and factors themselves; OR
 - 2. Full independence between breaks and factors, but limited dependence between breaks across series; OR
 - 3. Perfectly correlated breaks across series, but limited number of series undergoing breaks.
- Specifically, assume $|\Delta_i|$ for i = 1, ..., N is bounded; and
- at most $O(N^{1/2})$ series undergo a break.

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For "moderate" sized breaks, we can treat the instabilities as simply another additive error term which satisfies the usual assumptions.

$$X_t = \Lambda_t F_t + e_t = (\Lambda_0 + \xi_t) F_t + e_t = \Lambda_0 F_t + e_t + \xi_t F_t$$

Assume a one-factor model with Λ_0 known. Consider $\hat{F}_t = (\Lambda_0' \Lambda_0)^{-1} \Lambda_0' X_t$.

$$\hat{F}_t - F_t = (\Lambda_0'\Lambda_0)^{-1}\Lambda_0' e_t + (\Lambda_0'\Lambda_0)^{-1}\Lambda_0' \xi_t F_t$$
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- When $t > \kappa$, we have $\Lambda_0' \xi_t / N = (\lambda_{10} \Delta_1 + ... + \lambda_{N0} \Delta_N) / N$.
- Thus, this term is $O_p(B/N)$ where B is the number of non-zero Δ_i terms in the summation.

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Consistency under structural instability

As above, define B to be the number of series undergoing a break. Then:

$$\|\hat{F}_t - H'F_t\| = O_p\left(\max\left\{\frac{1}{\sqrt{N}}, \frac{1}{\sqrt{T}}\right\}, \left(\frac{B/\sqrt{N}}{\sqrt{N}}\right)\right\}\right)$$

- To summarise, the PCA estimator can accomodate a certain degree of temporal instability in the factor loadings.
- Under the right conditions, effects of instabilities can be eliminated asymptotically by averaging across series.

Aye!

"The [simulation] results [calibrated to previous empirical work by Stock and Watson] confirm...

Bates et al. (JoE 2013)

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- Bates et al. (JoE 2013)
- This is where the disagreement arises.

NAY!

"...in empirical applications parameters may change dramatically due to important economic events [...]."

Breitung & Eickmeier (JoE 2011)

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"...in empirical applications parameters may change dramatically due to important economic events [...]."

Breitung & Eickmeier (JoE 2011)

"...we find that a significant portion [around 80] of 132 U.S. macroeconomic time series have structural changes in their factor loadings. Although traditional principal components provide eight or more factors, there are significantly fewer nonspurious factors."

Yamamoto (J o Bus. & Econ. Stats. 2016)

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"...in empirical applications parameters may change dramatically due to important economic events [...]."

Breitung & Eickmeier (JoE 2011)

"...we find that a significant portion [around 80] of 132 U.S. macroeconomic time series have structural changes in their factor loadings. Although traditional principal components provide eight or more factors, there are significantly fewer nonspurious factors."

- Yamamoto (J o Bus. & Econ. Stats. 2016)
- Burgeoning literature on this side of the debate!

		IV. Baltagi et al. (2015)	

IV. Accounting for instabilities

- 47 ▶

Specifically, let us consider Baltagi, Kao & Wang (Manuscript, 2015)

- Paper is recent and its representation is fairly general; models in many other papers correspond to special cases.
- This paper tackles estimation of an approximate static factor model with a single abrupt break.
- Allows for changes in the number of factors and/or partial changes in factor loadings.

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Model description for X (in $T \times N$ matrix form)

$$X = \begin{bmatrix} F_{bef}^{0} \Lambda^{0'} + F_{bef}^{1} \Lambda^{1'} \\ F_{aft}^{0} \Lambda^{0'} + F_{aft}^{1} \Lambda^{2'} \end{bmatrix} + E$$

Notation:

► pre-change
$$\begin{cases} F_{bef}^0 = [f_1^0, ..., f_{\kappa}^0]'; & \kappa \times (r-q) \\ F_{bef}^1 = [f_1^1, ..., f_{\kappa}^1]'; & \kappa \times q \end{cases}$$

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Notation:

Image: A mathematical states and a mathem

Model description for X (in $T \times N$ matrix form)

$$X = \begin{bmatrix} F_{bef}^{0} \Lambda^{0'} + F_{bef}^{1} \Lambda^{1'} \\ F_{aft}^{0} \Lambda^{0'} + F_{aft}^{1} \Lambda^{2'} \end{bmatrix} + E$$

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Notation:

pre-change
$$\begin{cases} F_{bef}^{0} = [f_{1}^{0}, ..., f_{\kappa}^{0}]'; & \kappa \times (r-q) \\ F_{bef}^{1} = [f_{1}^{1}, ..., f_{\kappa}^{1}]'; & \kappa \times q \end{cases}$$
 post-change
$$\begin{cases} F_{aft}^{0} = [f_{\kappa+1}^{0}, ..., f_{T}^{0}]'; & (T-\kappa) \times (r-q) \\ F_{aft}^{1} = [f_{\kappa+1}^{1}, ..., f_{T}^{1}]'; & (T-\kappa) \times q \end{cases}$$
 factor loadings
$$\begin{cases} \Lambda^{0}; & N \times (r-q) \\ \Lambda^{1} \& \Lambda^{2}; & N \times q \end{cases}$$
 define
$$\begin{cases} \Lambda_{bef} = \left[\Lambda^{0} \quad \Lambda^{1}\right] \& \Lambda_{aft} = \left[\Lambda^{0} \quad \Lambda^{2}\right]; & N \times r \end{cases}$$

$$X = \begin{bmatrix} F_{bef}^{0} \Lambda^{0'} + F_{bef}^{1} \Lambda^{1'} \\ F_{aft}^{0} \Lambda^{0'} + F_{aft}^{1} \Lambda^{2'} \end{bmatrix} + E$$

Factor loadings and/or number of factors may change:

Structural change in the number of factors is incorporated as special case of structural change in factor loadings by allowing either Λ_{bef} or Λ_{aft} to have less than full column rank.

•
$$r = max\{r_{bef}, r_{aft}\}$$

- Emerging factors: If $r_1 < r_2$, some columns in Λ_{bef} are zeros.
- Disappearing factors: If $r_1 > r_2$, some columns in Λ_{aft} are zeros.
- For now, assume both Λ_{bef} and Λ_{aft} have full column rank.

A (10) > 4

Equivalent model with stable loadings

$$X = G\Gamma' + E$$

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Equivalent model with stable loadings

$$X = G\Gamma' + E$$

How do we define G and Γ ?

$$G_{T\times(r+q)} = \begin{bmatrix} \begin{bmatrix} F_{bef}^0 & F_{bef}^1 \end{bmatrix} A' \\ \begin{bmatrix} F_{aft}^0 & F_{aft}^1 \end{bmatrix} B' \end{bmatrix}$$

where

and

$$\begin{aligned} A_{(r+q)\times r} &= \begin{bmatrix} I_{(r-q)} & 0_{(r-q)\times q} \\ 0_{q\times (r-q)} & I_{q} \\ 0_{q\times (r-q)} & 0_{q} \end{bmatrix}, \\ B_{(r+q)\times r} &= \begin{bmatrix} I_{(r-q)} & 0_{(r-q)\times q} \\ 0_{q\times (r-q)} & 0_{q} \\ 0_{q\times (r-q)} & I_{q} \end{bmatrix}, \\ \Gamma_{N\times (r+q)} &= \begin{bmatrix} \Lambda^{0} & \Lambda^{1} & \Lambda^{2} \end{bmatrix}. \end{aligned}$$

Image: Image:



Suggested procedure

- 1. Obtain preliminary estimate, \hat{r} of number of factors, r, ignoring structural change.
 - Use standard methods developed for stable models.
 - $\blacktriangleright \lim_{(N,T)\to\infty} P(\hat{r}=r+q)=1.$

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Suggested procedure

- 1. Obtain preliminary estimate, \hat{r} of number of factors, r, ignoring structural change.
 - Use standard methods developed for stable models.
 - $\blacktriangleright \lim_{(N,T)\to\infty} P(\hat{r}=r+q)=1.$
- 2. Given \hat{r} , estimate the space spanned by the pseudo-factors, G_t .
 - Use PCA.

Image: A math a math

Suggested procedure

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- 2. Given \hat{r} , estimate the space spanned by the pseudo-factors, G_t .
 - Use PCA.
- 3. Estimate changepoint of factor loadings, κ .
 - Look for break in covariance matrix of pseudo-factors, Σ_G.
 - Note that $\Sigma_{G,bef} = A \Sigma_F A'$ and $\Sigma_{G,aft} = B \Sigma_F B'$.
 - Avoid potential infinite-dimensionality problem.

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Image: A math a math

- ▶ Split data into 2 subsamples for each "candidate" location of $\kappa \in [1, T 1]$.
- Pre-break subsample: $\hat{\Sigma}_{G,bef} = \sum_{t=1}^{\kappa} \hat{G}_t \hat{G}_t' / \kappa$
- Post-break subsample: $\hat{\Sigma}_{G,aft} = \sum_{t=\kappa+1}^{T} \hat{G}_t \hat{G}_t' / (T \kappa)$

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- Post-break subsample: $\hat{\Sigma}_{G,aft} = \sum_{t=\kappa+1}^{T} \hat{G}_t \hat{G}_t' / (T-\kappa)$
- Then, define sum of squared residuals:

$$\begin{split} \hat{S}(\kappa) &= \sum_{t=1}^{\kappa} [\textit{vec}(\hat{G}_t \hat{G}_t^{\ \prime} - \hat{\Sigma}_{G,bef})]'[\textit{vec}(\hat{G}_t \hat{G}_t^{\ \prime} - \hat{\Sigma}_{G,bef})] + \\ &\sum_{t=\kappa+1}^{T} [\textit{vec}(\hat{G}_t \hat{G}_t^{\ \prime} - \hat{\Sigma}_{G,aft})]'[\textit{vec}(\hat{G}_t \hat{G}_t^{\ \prime} - \hat{\Sigma}_{G,aft})] \end{split}$$

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- ▶ Split data into 2 subsamples for each "candidate" location of $\kappa \in [1, T 1]$.
- Pre-break subsample: $\hat{\Sigma}_{G,bef} = \sum_{t=1}^{\kappa} \hat{G}_t \hat{G}_t' / \kappa$
- Post-break subsample: $\hat{\Sigma}_{G,aft} = \sum_{t=\kappa+1}^{T} \hat{G}_t \hat{G}_t' / (T-\kappa)$
- Then, define sum of squared residuals:

$$\begin{split} \hat{S}(\kappa) &= \sum_{t=1}^{\kappa} [\textit{vec}(\hat{G}_t \hat{G}_t^{\ \prime} - \hat{\Sigma}_{G,bef})]'[\textit{vec}(\hat{G}_t \hat{G}_t^{\ \prime} - \hat{\Sigma}_{G,bef})] + \\ &\sum_{t=\kappa+1}^{T} [\textit{vec}(\hat{G}_t \hat{G}_t^{\ \prime} - \hat{\Sigma}_{G,aft})]'[\textit{vec}(\hat{G}_t \hat{G}_t^{\ \prime} - \hat{\Sigma}_{G,aft})] \end{split}$$

• Obtain LS estimator of changepoint: $\hat{\kappa} = \arg \min \hat{S}(\kappa)$.

Image: A math a math

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- Pre-break subsample: $\hat{\Sigma}_{G,bef} = \sum_{t=1}^{\kappa} \hat{G}_t \hat{G}_t' / \kappa$
- Post-break subsample: $\hat{\Sigma}_{G,aft} = \sum_{t=\kappa+1}^{T} \hat{G}_t \hat{G}_t' / (T \kappa)$
- Then, define sum of squared residuals:

$$\begin{split} \hat{S}(\kappa) &= \sum_{t=1}^{\kappa} [\operatorname{vec}(\hat{G}_t \, \hat{G}_t^{\ \prime} - \hat{\Sigma}_{G, bef})]' [\operatorname{vec}(\hat{G}_t \, \hat{G}_t^{\ \prime} - \hat{\Sigma}_{G, bef})] + \\ &\sum_{t=\kappa+1}^{T} [\operatorname{vec}(\hat{G}_t \, \hat{G}_t^{\ \prime} - \hat{\Sigma}_{G, aft})]' [\operatorname{vec}(\hat{G}_t \, \hat{G}_t^{\ \prime} - \hat{\Sigma}_{G, aft})] \end{split}$$

- Obtain LS estimator of changepoint: $\hat{\kappa} = \arg \min \hat{S}(\kappa)$.
- Estimator at most only "slightly misspecified": $\hat{\kappa} \kappa = O_p(1)$:

Image: A math a math

Determining the number of factors

- 4. Identify the number of factors based on $\hat{\kappa}$.
 - Split the data into 2 subsamples based on $\hat{\kappa}$.
 - Obtain \hat{r}_{bef} and \hat{r}_{aft} using standard criteria.
 - \hat{r}_{bef} and \hat{r}_{aft} shown to be robust to "slightly misspecified" $\hat{\kappa}$:

< 17 ▶

Estimating the factor space

- 5. Estimate the factor space based on $\hat{\kappa}$ and \hat{r}_{bef} , \hat{r}_{aft} .
 - Use PCA on each of the subsamples.
 - Convergence rate is the same as in Bai & Ng (Econometrica 2002).

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		V. My research

V. Sequential detection of structural instabilities in DFMs

ended model

V. Bates et al. (2013)

Do structural instabilities matter?

NAY!

- Stock & Watson (JASA 2002)
- Bates, Plagborg-Moller, Stock & Watson (JoE 2013)

Aye!

- Breitung & Eickmeier (JoE 2011)
- Corradi & Swanson (Forthcoming, JoE)
- Chen, Dolado & Gonzalo (JoE 2014)
- Han & Inoue ('metric Theory 2014)
- Baltagi, Kao & Wang (Manuscript, 2015)
- Yamamoto & Tanaka (JoE 2015)
- Sabharwal (Journal of Optimism, 2016)

Thank you for your attention!

Contact

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